

Building Random Trees from Blocks

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Joint work with

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The Probability Model

- We have a finite collection of unlabeled, rooted, nonplanar building blocks (trees) $\mathcal{C} = \{T_1, \dots, T_k\}$, that occur with respective probabilities p_1, \dots, p_k ; $\sum_{j=1}^k p_j = 1$.

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- A special case: \mathcal{C} consists of only one node; the block tree is isomorphic to the well-studied standard recursive tree.

Example



Figure : A collection of building blocks of size 4, with probabilities $\frac{1}{3}$, and $\frac{2}{3}$, respectively.

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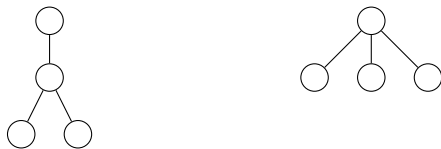


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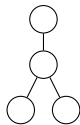


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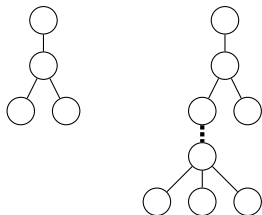


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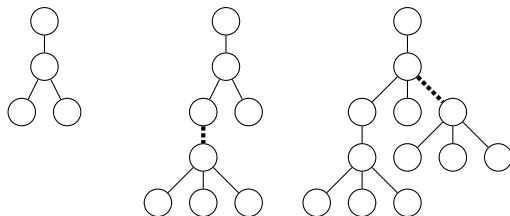


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Step 3: The third block occurs with probability $\frac{2}{3}$, and the parent is chosen with probability $\frac{1}{8}$.

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 - Height of the blocks tree

Leaves

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A leaf is a node that has no children.

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Theorem

Let L_n be the number of leaves in a random tree built from the building blocks T_1, \dots, T_k , which are selected at each step with probabilities p_1, \dots, p_k . Let $\mathbf{E}[\Lambda_{\mathcal{C}}]$ be the average number of leaves in the given collection, and $\mathbf{Var}[\Lambda_{\mathcal{C}}]$ be the variance. Then,

$$\frac{L_n - \frac{t \mathbf{E}[\Lambda_{\mathcal{C}}]}{t+1} n}{\sqrt{n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \sigma_{\mathcal{C}}^2),$$

where

$$\sigma_{\mathcal{C}}^2 := \left(\frac{\mathbf{Var}[\Lambda_{\mathcal{C}}]}{t+2} + \frac{\mathbf{E}[\Lambda_{\mathcal{C}}](t+1 - \mathbf{E}[\Lambda_{\mathcal{C}}])}{(1+t)^2(2+t)} \right) t.$$

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- This will enable us to use the powerful theory of Pólya urns.
- Suppose T_i has ℓ_i leaves (and consequently it has $t - \ell_i$ internal nodes).
- Let $\Lambda_{\mathcal{C}}$ be number of leaves in a randomly chosen block, i.e., $\Lambda_{\mathcal{C}}$ has probability mass

$$P(\Lambda_{\mathcal{C}} = \ell) = \sum_j p_j,$$

where the sum is taken over all j such that block T_j has ℓ leaves.

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The replacement matrix for this urn:

$$\mathbf{A} = \left(\begin{array}{c|cc} & W & B \\ \hline W & \Lambda_{\mathcal{C}} - 1 & t - \Lambda_{\mathcal{C}} + 1 \\ B & \Lambda_{\mathcal{C}} & t - \Lambda_{\mathcal{C}} \end{array} \right)$$

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Row sum for this matrix is a constant! Balanced Urns!

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For balanced urns it is shown by [Athreya 1968] that

$$\frac{W_n}{n} \xrightarrow{\text{a.s.}} \lambda_1 v_1,$$

where λ_1 is the principal (largest real) eigenvalue of the average of the replacement matrix, and (v_1, v_2) is the corresponding left eigenvector of $\mathbf{E}[\mathbf{A}]$.

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and hence,

$$\frac{W_n}{n} \xrightarrow{\text{a.s.}} \frac{t}{t+1} \mathbf{E}[\Lambda_{\mathcal{L}}].$$

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Further we also notice that $\Re\lambda_2 < \frac{1}{2}\lambda_1$.

[Smythe 1996] showed that

$$\frac{W_n - \lambda_1 v_1}{\sqrt{n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \sigma^2),$$

for some variance σ^2 and states that σ^2 is generally hard to compute, but we will obtain the exact variance of W_n .

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Hence the theorem follows!

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Let D_n be the depth of the root of the n th inserted block in a random tree built from the building blocks T_1, \dots, T_k , which are selected at each step with probabilities p_1, \dots, p_k . Let $\mathbf{E}[\Delta_{\mathcal{C}}]$ be the average depth of a node in the given collection, and $\mathbf{Var}[\Delta_{\mathcal{C}}]$ be the variance of that depth. Then,

$$\frac{D_n - (\mathbf{E}[\Delta_{\mathcal{C}}] + 1) \ln n}{\sqrt{\ln n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \mathbf{Var}[\Delta_{\mathcal{C}}] + (\mathbf{E}[\Delta_{\mathcal{C}}] + 1)^2).$$

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- $\Delta_{\mathcal{C}}$ (representing the depth of a parent node), which is completely determined by the structure of the trees in the collection;
- Each δ_n has the same random distribution as $\Delta_{\mathcal{C}}$.

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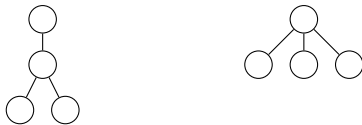


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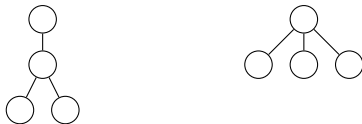


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$$\begin{aligned} \mathbf{E}[e^{D_n u} | \mathcal{F}_{n-1}] &= \mathbf{E}\left[\sum_{i=1}^{n-1} e^{(D_i + \delta_n + 1)u} \frac{1}{n-1} \middle| \mathcal{F}_{n-1}\right] \\ &= \frac{1}{n-1} \mathbf{E}[e^{(\delta_n + 1)u}] \sum_{i=1}^{n-1} e^{D_i u}. \end{aligned}$$

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Solving this full history recurrence and after appropriate centering and scaling we get

$$\mathbf{E} \left[\exp \left(\left(\frac{D_n - (\mathbf{E}[\Delta_{\mathcal{C}}] + 1) \ln n}{\sqrt{\ln n}} \right) u \right) \right] \rightarrow e^{\frac{1}{2} (\mathbf{Var}[\Delta_{\mathcal{C}}] + (\mathbf{E}[\Delta_{\mathcal{C}}] + 1)^2) u^2}.$$

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Therefore we have the theorem,

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$$\phi_{D_n}(u) = \frac{e^u \psi_{\mathcal{C}}(u)}{n-1} \sum_{i=1}^{n-1} \phi_{D_i}(u),$$

valid for $n \geq 2$, with the initial condition $\phi_{D_1}(u) = 1$.

Solving this full history recurrence and after appropriate centering and scaling we get

$$\mathbf{E} \left[\exp \left(\left(\frac{D_n - (\mathbf{E}[\Delta_{\mathcal{C}}] + 1) \ln n}{\sqrt{\ln n}} \right) u \right) \right] \rightarrow e^{\frac{1}{2} (\mathbf{Var}[\Delta_{\mathcal{C}}] + (\mathbf{E}[\Delta_{\mathcal{C}}] + 1)^2) u^2}.$$

Therefore we have the theorem,

$$\frac{D_n - (\mathbf{E}[\Delta_{\mathcal{C}}] + 1) \ln n}{\sqrt{\ln n}} \xrightarrow{\mathcal{D}} \mathcal{N}(0, \mathbf{Var}[\Delta_{\mathcal{C}}] + (\mathbf{E}[\Delta_{\mathcal{C}}] + 1)^2).$$

Remark. The expressions for mean and variance are valid, even if $t = t(n)$ grows with n . However, in the asymptotic derivations of the central limit theorem we have to keep t relatively small, compared to n .

Total path length

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Theorem

Let X_n be the total path length of a tree built from the blocks of a collection \mathcal{C} . Then, there is an absolutely integrable random variable X , such that $X_n/n - (t + \mathbf{E}[\chi_{\mathcal{C}}]) H_n + t$ converges to X , both in \mathcal{L}_2 and almost surely.

Sketch of Proof.

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- For example:



Figure : A collection of building blocks of size 4, with probabilities $\frac{1}{3}$, and $\frac{2}{3}$, respectively.

$$P(\chi_{\mathcal{C}} = 5) = \frac{1}{3}, \quad \text{and} \quad P(\chi_{\mathcal{C}} = 3) = \frac{2}{3}.$$

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- If the n th block is adjoined to a node $v \in \mathcal{T}_{n-1}$, at depth $\tilde{D}(v)$ in the tree \mathcal{T}_{n-1} then each node in the last inserted block appears at distance equal to $\tilde{D}(v) + 1$, plus its own depth in the last block.

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- So we have the following stochastic recurrence for $n \geq 2$:

$$\begin{aligned}\mathbf{E}[X_n | \mathcal{F}_{n-1}] &= X_{n-1} + t \left(\frac{1}{t(n-1)} \sum_{v \in \mathcal{T}_{n-1}} \tilde{D}(v) + 1 \right) + \mathbf{E}[\chi_{\mathcal{C}} | \mathcal{F}_{n-1}] \\ &= X_{n-1} + \frac{X_{n-1}}{n-1} + t + \mathbf{E}[\chi_{\mathcal{C}}].\end{aligned}$$

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- Solving we get

$$\mathbf{E}[X_n] = (t + \mathbf{E}[\chi_{\mathcal{C}}]) n H_n - nt \sim (t + \mathbf{E}[\chi_{\mathcal{C}}]) n \ln n,$$

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- Squaring and taking double expectation we get,

$$\begin{aligned} \mathbf{E}[X_n^2] &= \frac{n+1}{n-1} \mathbf{E}[X_{n-1}^2] + \frac{2n}{n-1} (\mathbf{E}[\chi_{\mathcal{C}}] + t) \mathbf{E}[X_{n-1}] \\ &\quad + t^2 \mathbf{E}[\check{D}_n^2] + t^2 + 2t \mathbf{E}[\chi_{\mathcal{C}}] + \mathbf{E}[\chi_{\mathcal{C}}^2]. \end{aligned}$$

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- We develop a separate recurrence for $\mathbf{E}[\check{D}_n^2]$ and solving for $\mathbf{E}[X_n^2]$ we get,

$$\begin{aligned} \mathbf{Var}[X_n] &\sim \left(t^2 (\mathbf{E}[\Delta_{\mathcal{C}}^2] + 2(\mathbf{E}[\Delta_{\mathcal{C}}])^2 + 4\mathbf{E}[\Delta_{\mathcal{C}}] + 4) + \mathbf{E}[\chi_{\mathcal{C}}^2] + 4t \mathbf{E}[\chi_{\mathcal{C}}] \right. \\ &\quad \left. - (\mathbf{E}[\chi_{\mathcal{C}}] + t)^2 \left(2 + \frac{\pi^2}{6} \right) \right) n^2. \end{aligned}$$

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- We know the conditional relation:

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Remark. The expressions for mean and variance are valid, even when $t = t(n)$ is no longer fixed but grows with n .

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Theorem

Let H_n be the height of a random tree built from the building blocks T_1, \dots, T_k , which are selected at each step with probabilities p_1, \dots, p_k . We then have

$$\frac{H_n}{\ln n} \xrightarrow{\text{a.s.}} e(\mathbf{E}[\Delta_{\mathcal{C}}] + 1).$$

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Hence

$$H_n \geq (1 + \hat{\delta}_1) + (1 + \hat{\delta}_2) + \dots + (1 + \hat{\delta}_{\hat{H}_n}) = \hat{H}_n + \sum_{i=1}^{\hat{H}_n} \hat{\delta}_i,$$

where $\hat{\delta}_i$, for $i = 1, \dots, \hat{H}_n$ are all independent.

Sketch of Proof.

- By the strong law of large numbers $\frac{1}{\hat{H}_n} \sum_{i=1}^{\hat{H}_n} \hat{\delta}_i \xrightarrow{\text{a.s.}} \mathbf{E}[\Delta_{\mathcal{C}}]$.

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- Then, the height of the blocks tree is bounded above:

$$H_n \leq \max_{1 \leq i \leq n} \left\{ (1 + \hat{\delta}_1^{(i)}) + (1 + \hat{\delta}_2^{(i)}) + \cdots + (1 + \hat{\delta}_{\hat{D}_i}^{(i)}) \right\}$$

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where $\hat{\delta}_{\hat{D}_i+1}^{(i)}, \dots, \hat{\delta}_{\hat{H}_n}^{(i)}$ are additional independent random variable padded at the end to make all the expressions of the same length.

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Combining the bounds the result follows!

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- The distribution of the number of nodes of outdegree $j > 0$.

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- We can not handle a collection of blocks of different sizes, which would be more general.
- The underlying structure here is a recursive tree. We could try to build other types of tree from blocks.

Thank You!
Questions are welcome!